

Are Frequency and Severity independent in health insurance?

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11.02.2021

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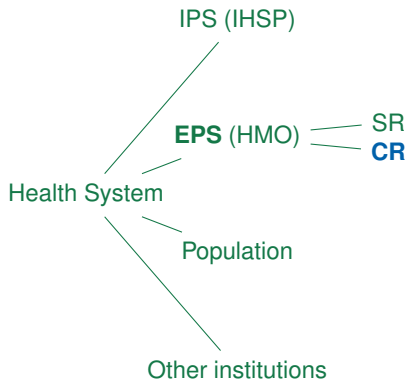
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Context ...

System [1–7]



- IHSP: Institutional Health Service Providers.
- HMO: Health Maintenance Organization.

Contributory Regime - CR [1, 8–12]

VHI (Commercial Premiums)

Compound Processes

VHI (CRM) [13]

Separated Processes

Separated Processes [13–15]

- Claim.
- Exposure.
- Pure Premium.
- **Frequency.**
- **Severity.**

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Motivation

> **Factors:** Exposure \rightarrow Frequency \leftrightarrow Severity \leftarrow Exposure.

Motivation

> Frequency - Severity Independence (FSI) assumption:

- It is “*frequently adopted in the non-life insurance technique*” [13] and it is an *a priori* condition [16, 17].
- It provides some advantages: Statistical inference, less computationally-intensive processes, ... [17–20].
- It is restrictive and invalid [17, 21, 22]
- It could negatively impact on:
 - Total loss (underestimation [9, 22] or overestimation [21]).
 - Commercial premiums (*Rate making*) [16].
 - Risk control and risk management [17, 23].
- The correct modelling is relevant in public health: Systemic risk and insurance issues [23].

Goal

Objective

We aim to provide evidence of the relationship between the claim frequency and claim severity by testing the independence assumption hypothesis using the Multivariate Approach, which would also derives the dependence structure.

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Model

$$S = \begin{cases} 0 & N = 0 \\ \sum_{i=1}^N X_i & N > 0 \end{cases} \quad (1)$$

- $N \in \mathbb{Z}^* \mid \mathbb{Z}^* = \{0\} \cup \mathbb{Z}^+$
- $Y_j \in \mathbb{R}^* \mid \mathbb{R}^* = \{0\} \cup \mathbb{R}^+$

- $X_j = \Pi(Y_j)$ tal que $X_j \leq Y_j$.

Univariate Approach

FSM-IUA

$$\mathbb{E}[S] = \mathbb{E}[X_1] \mathbb{E}[N]$$

[13, 14, 24]

FSM-CUA

$$f(N, X) = f(N) \times f(X|N)$$

[16, 17, 19, 20]

Multivariate Approach

FSM-MA

$$f(X, N) = h(X, N) \leftarrow H(X, N) \text{ [21, 22, 25]}$$

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Test

FSM - CUA

$$f(N, X; \Theta) = f(N; \theta_1) \times f(X | N; \theta_2)$$

$$\theta_2 = \{\theta^{X|N}, \theta_X\}$$

- $\hat{\theta}_{X|N} = 0$
- $\hat{\theta}_{X|N} > 0$
- $\hat{\theta}_{X|N} < 0$

FSM - MA

$$f(N, X; \Theta) = H(N, X; \gamma)$$

- $\hat{\gamma} = 0$
- Association measures: $\delta(\cdot)$

Literature

Non-health insurance

- Collision insurance: small $\hat{\theta}_{X|N} < 0$ [17].
- Liability insurance: $\hat{\theta}_{X|N} > 0$ [20].
- Car insurance:
 - $\hat{\theta}_{X|N} > 0$, $\hat{\delta} > 0$, $\hat{\gamma} \neq 0$ [16].
 - $\hat{\delta} > 0$, $\hat{\gamma} \neq 0$ - Underestimation [22].
 - $\hat{\delta} < 0$, $\hat{\gamma} \neq 0$ - Overestimation [21].

Health insurance

- Prediction models for expenditure [26]:
 - Inpatient: $\hat{\theta}_{X|N} = 0$.
 - Outpatient: $\hat{\theta}_{X|N} > 0$.
- Bivariate distributions of drug claims and other health claims: $\hat{\gamma} \neq 0$ [27].
- Health insurance and health care demand (married couples) [28].
- CRM with dependencies between Frequency and Severity with different hospital settings [29].

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Definition

- F - CA: Frequency (N_i)- Claim Amount (X_i) [13, 17].
- F - ACA: Frequency (N_i) - Average Claim Amount (X_i/N_i) [10, 12, 22].
- Exposure: Sum of affiliates weighted by their period of membership.

Method

Independence test

Independence $\Rightarrow (\delta = 0)$

$(\delta \neq 0) \Rightarrow$ Dependence

$$H_o : \delta = 0$$

Dependence Structure

Measures of association [30–36]

- Pearson: r_P .
 - Linearity.
 - Normality.
- Spearman: ρ
- Kendall: τ

Copula [31, 33, 34, 37–44]

- 1 Copula specification.
 - Parametric estimation: IFME (*Inference Functions for Margins Estimator*).
 - Non-parametric estimation: MPLE (*Maximum Pseudo-Likelihood Estimator*).
- 2 Measures of association with transformation of the copula parameter:
 - Kendall: τ
 - Blomqvist: β
 - Lower Tail: λ_L
 - Upper Tail: λ_U

Copula

Theorem (Sklar's theorem)

$$H(x, y) = C(F(x), G(y))$$

where H is a joint distribution function with margins F and G for random variables X and Y , respectively, the copula C is unique if F and G are continuous. Although the claim frequency marginal is discrete, which means that C is uniquely determined on $\text{Ran}F \times \text{Ran}G$ [33, Theorem 2.3.3].

However, the copula representation and its non-uniqueness do not “(...) inhibit empirical applications” [40]

Copula Estimation and Association

IFME

- 1 Theoretical distributions [41]
 - Severity: *Gamma, Log-Normal, Exponential, Log-Logistic (Fisk), Pareto and Gumbel.*
 - Frequency: *Geometric, Poisson and Negative Binomial.*
- 2 MLE.

PMLE [38]

- 1 Pseudo-observations.
- 2 MLE.

Fréchet-Hoeffding bounds [31, 33, 43]

$$W(\mathbf{u}) \leq C(\mathbf{u}) \leq M(\mathbf{u})$$

- $C(\mathbf{u}) \rightarrow M(\mathbf{u})$: Comonotonic
 $\equiv \delta(N, X) = 1.$
- $C(\mathbf{u}) \rightarrow W(\mathbf{u})$: Countermonotonic
 $\equiv \delta(N, X) = -1.$

Copulas [33, 38]

- Gaussian, Clayton, Frank, Gumbel, Joe, BB6 and BB8.
- Selection by AIC and Kendall's method (GoF: White and Kendall) [38].

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Data

> **Data:** HMO - VHI (EPS-PAC): 35,744 affiliates.

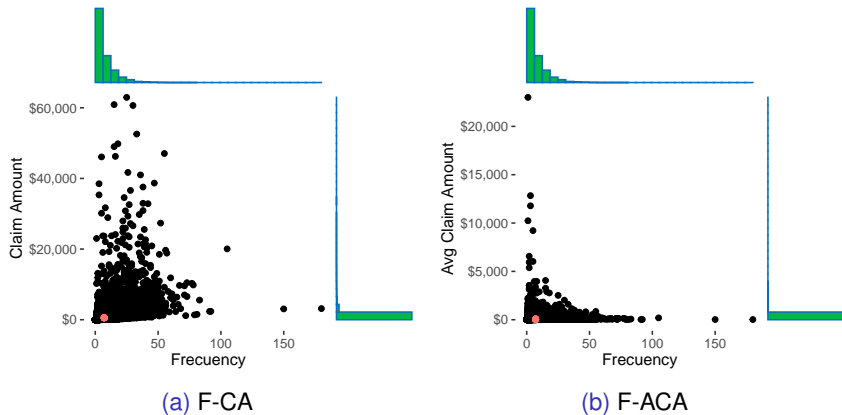
Table: Descriptive statistics

| Variables | All | Female | Male | Over 60 y/o | Under 60 y/o |
|-------------------------|-----------------|-----------------|----------------|--------------------|---------------------|
| Claim Amount | \$ 19,788,384.7 | \$ 12,281,185.0 | \$ 7,507,199.8 | \$ 5,600,321.0 | \$ 14,188,063.8 |
| Number of Claims | 254,155 | 166,883 | 87,272 | 54,523 | 199,632 |
| Affiliates | 35,744 | 20,523 | 15,221 | 4,724 | 31,020 |
| Exposure | 32,275 | 18,562 | 13,713 | 4,536 | 27,739 |
| Severity | \$ 77.9 | \$ 73.6 | \$ 86 | \$ 102.7 | \$ 71.1 |
| Claim Amount per Person | \$ 553.6 | \$ 598.4 | \$ 493.2 | \$ 1,185.5 | \$ 457.4 |
| Frequency | 7.9 | 9.0 | 6.4 | 12.0 | 7.2 |
| Pure Premium | \$ 613.1 | \$ 661.6 | \$ 547.5 | \$ 1,234.6 | \$ 511.5 |

Source: Authors

Features

Figure: Marginal Plots of F-CA and F-ACA



Source: Authors

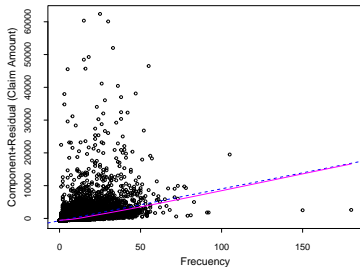
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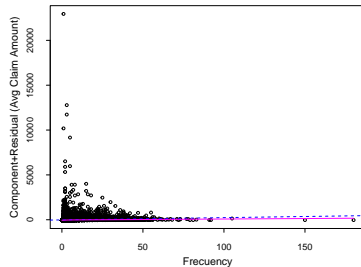
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Gaussian?



(c) F-CA



(d) F-ACA

Figure: Partial residuals plots of F-CA and F-ACA

Gaussian?

Table: Shapiro-Wilk test

| Group | Statistic | p - value |
|--------------|------------------|------------------|
| F-CA | 0.2184 | < 0.01 |
| F-ACA | 0.234 | < 0.01 |

Source: Authors

Measures of Association

Table: Measures of association

| Group | Method | Outliers | |
|-------|----------|-----------|-----------|
| | | Yes | No |
| F-CA | r_P | 0.4482*** | 0.6733*** |
| | ρ_S | 0.9102*** | 0.9099*** |
| | τ_K | 0.7732*** | 0.7726*** |
| F-ACA | r_P | 0.0934*** | 0.1615*** |
| | ρ_S | 0.6707*** | 0.665*** |
| | τ_K | 0.5115*** | 0.5059*** |

Source: Authors

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Table: Fitted theoretical distributions by IFME

| Variable | Distribution | Parameters | Estimate | Std.Error | p-value |
|----------------------|-------------------|---------------|----------|-----------|---------|
| Frequency | Negative Binomial | θ_{NB} | 0.803 | 0.007 | < 0.01 |
| | | μ_{NB} | 6.915 | 0.043 | < 0.01 |
| Claim Amount | Log-normal | μ_{LN} | 5.302 | 0.009 | < 0.01 |
| | | σ_{LN} | 1.539 | 0.006 | < 0.01 |
| Average Claim Amount | Log-logistic | β_{LI} | 1.946 | 0.010 | < 0.01 |
| | | α_{LI} | 34.856 | 0.184 | < 0.01 |

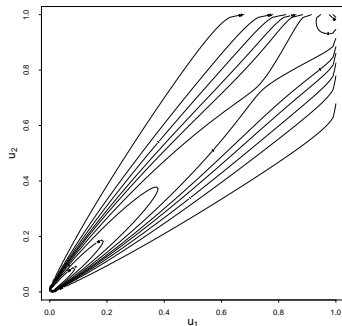
Source: Authors

Copula

- Parametric estimation: IFME

- Frequency: Negative Binomial.
- Severity 1: Log-Normal.
- Severity 2: Log-logistic (Fisk).

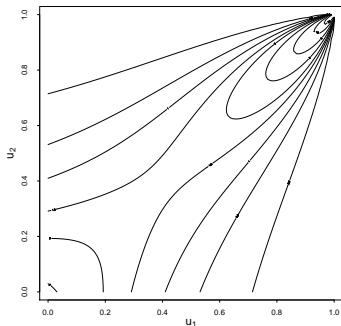
Survival BB6 ($\theta_1=5.764$, $\theta_2=1.37$)



- Non-parametric estimation: MPLE:

- Pseudo-observations.
- Graphics of F-CA Non-parametric.

Joe ($\theta=3.216$)



Copula and Association

Table: Association measures by Copula

| Class | Group | Method | Copula | Parameter 1 | Parameter 2 | τ_K | β_B | λ_L | λ_U |
|----------------|-------|---------|--------------|-------------|-------------|----------|-----------|-------------|-------------|
| Parametric | F-CA | AIC | Survival BB8 | 5.214 | 1 | 0.56 | 0.606 | 0 | 0 |
| | | Kendall | Joe | 2.672 | | 0.475 | 0.486 | 0 | 0.704 |
| | F-ACA | AIC | Survival BB8 | 6 | 0.438 | 0.314 | 0.35 | 0 | 0 |
| | | Kendall | Joe | 1.443 | | 0.2 | 0.191 | 0 | 0.383 |
| Non-parametric | F-CA | AIC | Survival BB6 | 5.764 | 1.37 | 0.791 | 0.819 | 0.908 | 0 |
| | | Kendall | Joe | 3.216 | | 0.543 | 0.561 | 0 | 0.759 |
| | F-ACA | AIC | Survival Joe | 3.867 | | 0.603 | 0.629 | 0.804 | 0 |
| | | Kendall | Joe | 1.613 | | 0.255 | 0.248 | 0 | 0.463 |

Source: Authors

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Limitations

- Copula: Gaussian, Clayton, Gumbel, Frank, Joe, BB8, and BB6.
- No Zero-Inflated Models.
- The VHI covers different hospital settings.
- The VHI is a “Plan de Atención Complementaria”: Dependence factors ← e.g. Hospital Network.

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Conclusion

What did we get?

- Rejection of the H_0 : FSI.
- No linearity - No bivariate normal.
- IFME underestimates the association.
- Positive association.
- Tail dependence.
- The structure: Joe Copula is remarkable.

What's next?

- > Following [22], if the association is positive, it could be an underestimation of the total loss in car insurance.
 - **Warning**: Does the FSI lead to a underestimation of total loss in health insurance?
 - Premium reserves.
 - VHI.
 - Compulsory Health Insurance and Social Health Insurance.

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Copula Regression

- **Conditional copulas** [34, 38]: *They capture the dependence among the components of \mathbf{X} conditionally on \mathbf{Z} [38].*
- If C has a density (mixed derivate of order d): $c \Rightarrow$ from the CDF $H_{\mathbf{Z}}(\mathbf{X}) = C_{\mathbf{z}}(F_{Z,1}(x_1), \dots, F_{Z,d}(x_d))$, the PDF is:

$$h_{\mathbf{z}} = c_{\varphi(\mathbf{z})}(F_{1,\varphi_1(\mathbf{z})}(x_1), \dots, F_{d,\varphi_d(\mathbf{z})}(x_d)) \prod_{j=1}^d f_{j,\varphi_j(\mathbf{z})}(x_j)$$

- Conditional log-likelihood function of the model is:

$$L_n(\beta, \beta_1, \dots, \beta_d) = \sum_{i=1}^n \log c_{\phi}(\mathbf{Z}_i) (F_{1,\phi_1(\mathbf{Z}_i)}(X_{i1}), \dots, F_{d,\phi_d(\mathbf{Z}_i)}(X_{id})) + \sum_{j=1}^d \sum_{i=1}^n \log f_{j,\phi_j(\mathbf{Z}_i)}(X_{ij})$$

Copula Regression: Traditional case

- $d = 2$ y $n = 35,744^*$

$$L_n(\beta, \beta_1, \beta_2) = \underbrace{\sum_{i=1}^n \log c_{\phi}(\mathbf{Z}_i) (F_{1, \phi_1}(\mathbf{Z}_i)(X_{i1}), F_{2, \phi_2}(\mathbf{Z}_i)(X_{i2}))}_{\text{JoeCopula}} + \underbrace{\sum_{i=1}^n \log f_{1, \phi_1}(\mathbf{Z}_i)(X_{i1})}_{\text{Poisson}} + \underbrace{\sum_{i=1}^n \log f_{2, \phi_2}(\mathbf{Z}_i)(X_{i2})}_{\text{Gamma}}$$

- Tweedie: *Compound Poisson-gamma process* [11, 14].
- FSM - IUA: Frequency (Poisson) - Severity (Gamma).
- FSM - MA: Frequency (Poisson) - Severity (Gamma).

Example

Table: Estimated Total Loss

| Pure Premium | Value | Std. Dev. | Lower Conf. Int | Upper Conf. Int | P-value |
|----------------------|---------|-----------|-----------------|-----------------|---------|
| Tweedie | 667.49 | 276.79 | 661.07 | 673.90 | < 0.001 |
| Ind. Poisson – Gamma | 631.49 | 324.56 | 623.97 | 639.02 | < 0.001 |
| Ind. NB – Gamma | 631.40 | 313.54 | 624.13 | 638.67 | < 0.001 |
| Cop. Poisson – Gamma | 631.28 | 324.78 | 623.75 | 638.81 | < 0.001 |
| Cop. NB – Gamma | 650.50* | 318.65 | 643.12 | 657.89 | < 0.001 |

Source: Authors

$$\frac{\hat{E}_{\text{independence}}(PP)}{\hat{E}_{\text{Joe}}(PP)} = 0.97^{**}$$

- * Greater than real (613.1).
- ** 0.934 with Clayton for car insurance [22] - **Preliminary Findings**
- Marginal misspecification \Rightarrow Fisk: Pure Premium (Ind)= USD 598.41.

¡Muchas gracias!

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